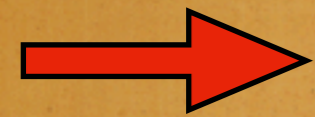


RECAP LECTURE 4 + FRITS PAERELS

THERMAL LIMIT OF STOCHASTIC RADIATION PROCESSES



$$h\nu \ll kT \quad \text{THERMAL LIMIT} \quad \overline{\Delta P^2}(\nu) = \bar{P}^2(\nu)$$

PAERELS SHOWED THAT ONE DOESN'T ALWAYS GET THE EXACT POISSON QUANTUM NOISE AT $h\nu \gg kT$



USE INCOMPLETE GAMMA FUNCTION TO CALCULATE POISSON AND GAUSS CUMULATIVE DISTRIBUTION FUNCTION



PROPAGATION OF ERRORS

UNDER THE ASSUMPTION OF INDEPENDENT VARIABLES

$$\bar{f} = f(\bar{u}, \bar{v}, \dots)$$

$$\sigma_f^2 = \sigma_u^2 \left(\frac{\partial f}{\partial u} \right)^2 + \sigma_v^2 \left(\frac{\partial f}{\partial v} \right)^2 + \dots$$

TODAY:

COMPARING DATA WITH A MODEL:
LEAST-SQUARES FITTING, MAXIMUM
LIKELIHOOD METHOD: GAUSSIAN DATA

“REAL” MAXIMUM LIKELIHOOD
METHOD: POISSONIAN DATA

OAF2 CHAPTER 5.3+5.4
SEE ALSO NUM RES CHAPTER 15

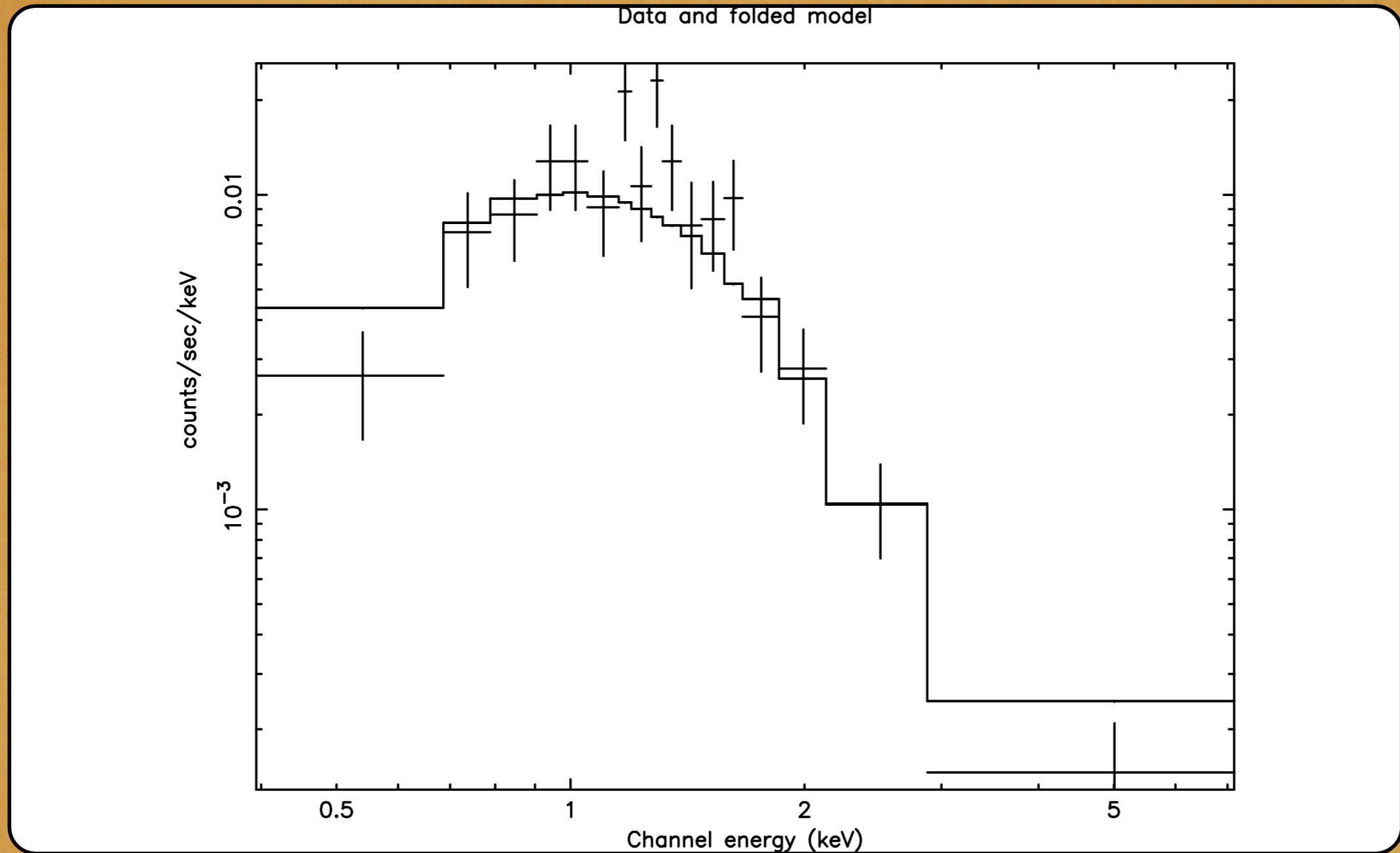
COMPARE DATA WITH A MODEL

DESCRIBE DATA IN TERMS OF A CONTINUOUS
FUNCTION

COMPARE OBSERVATIONS (DATA) WITH
THEORETICAL MODEL PREDICTION

DESCRIBE THE DATA IN A FEW PARAMETERS

EXAMPLE: CONTINUOUS MODEL THROUGH DISCRETE DATA & MODEL PREDICTION



X-RAY BINARY IN QUIESCENCE NEUTRON STAR
ATMOSPHERE MODEL

MAXIMUM LIKELIHOOD: MOST LIKELY
OUTCOME IS ASSUMED TO BE THE
'CORRECT' ONE

METHOD OF LEAST SQUARES

$$dQ_i = P_i dx$$

PROBABILITY DENSITY FUNCTION $\frac{dQ_i}{dx} = P_i$
↪ GAUSS, POISSON

$$P(y_i) \Delta y = \frac{1}{\sqrt{2\pi}\sigma_i} \exp\left(-\frac{1}{2} \frac{(y_i - y_m)^2}{\sigma_i^2}\right) \Delta y$$

NOTE: y_m = MODEL VALUE NOT MEAN HERE!

METHOD OF LEAST SQUARES

$$P \propto \prod_{i=1}^N \left\{ \exp \left[-\frac{1}{2} \left(\frac{y_i - y_m}{\sigma_i} \right)^2 \right] \right\}$$

$$\propto \exp \left[-\frac{1}{2} \sum_{i=1}^N \left(\frac{y_i - y_m}{\sigma_i} \right)^2 \right]$$

MINIMISE: $\chi^2 \equiv \sum_{i=1}^N \left(\frac{y_i - y_m}{\sigma_i} \right)^2$

MINIMISATION: ROOT FINDING PROBLEM

1 D: $\frac{\partial}{\partial y_i} \sum_{i=1}^N \left(\frac{y_i - y_m}{\sigma_i} \right)^2 = 0$

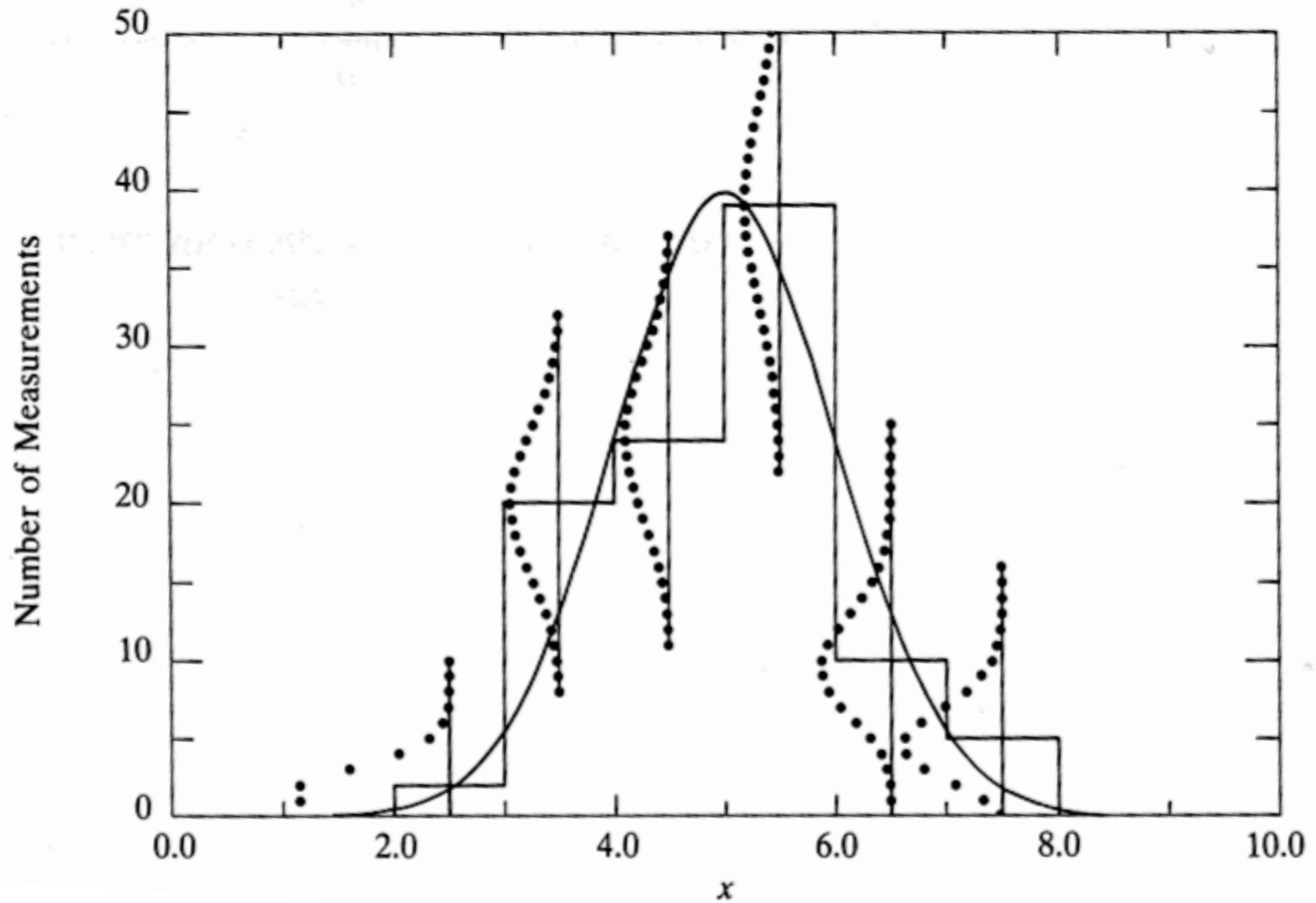
MORE ABOUT χ^2

DRAWN FROM NORMAL DISTRIBUTION
DISTRIBUTION OF χ_i^2 IS A χ^2 DISTRIBUTION
FOR N MEASUREMENTS DESCRIBED BY
M VARIABLES, THERE ARE N-M
DEGREES OF FREEDOM (D.O.F.)

PROBABILITY OF OBTAINING A
CERTAIN χ^2

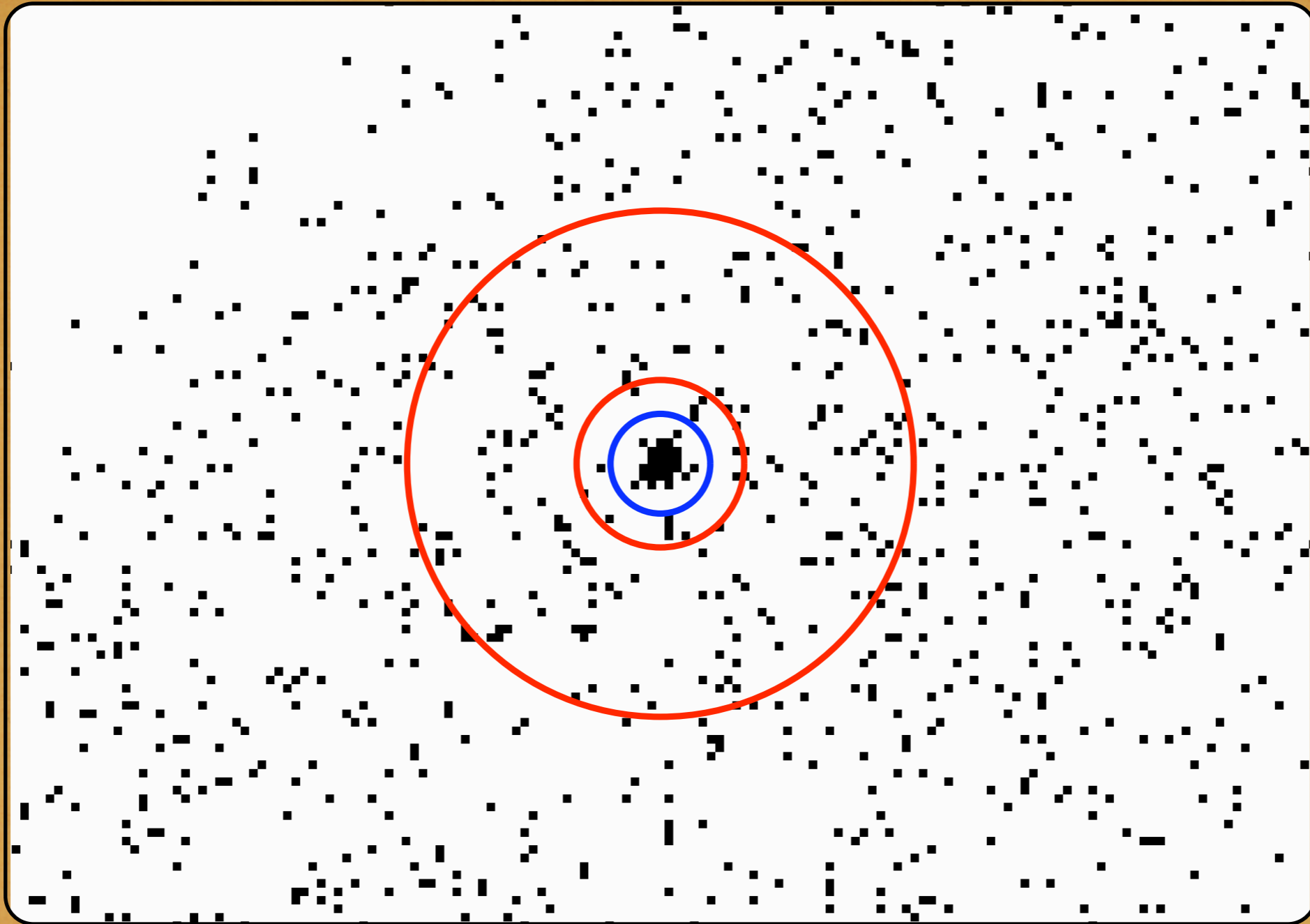
$$P(\chi_{obs}^2) = \text{gammq}\left(\frac{N - M}{2}, \frac{\chi_{obs}^2}{2}\right)$$

VALUE AND POISSON ERRORS

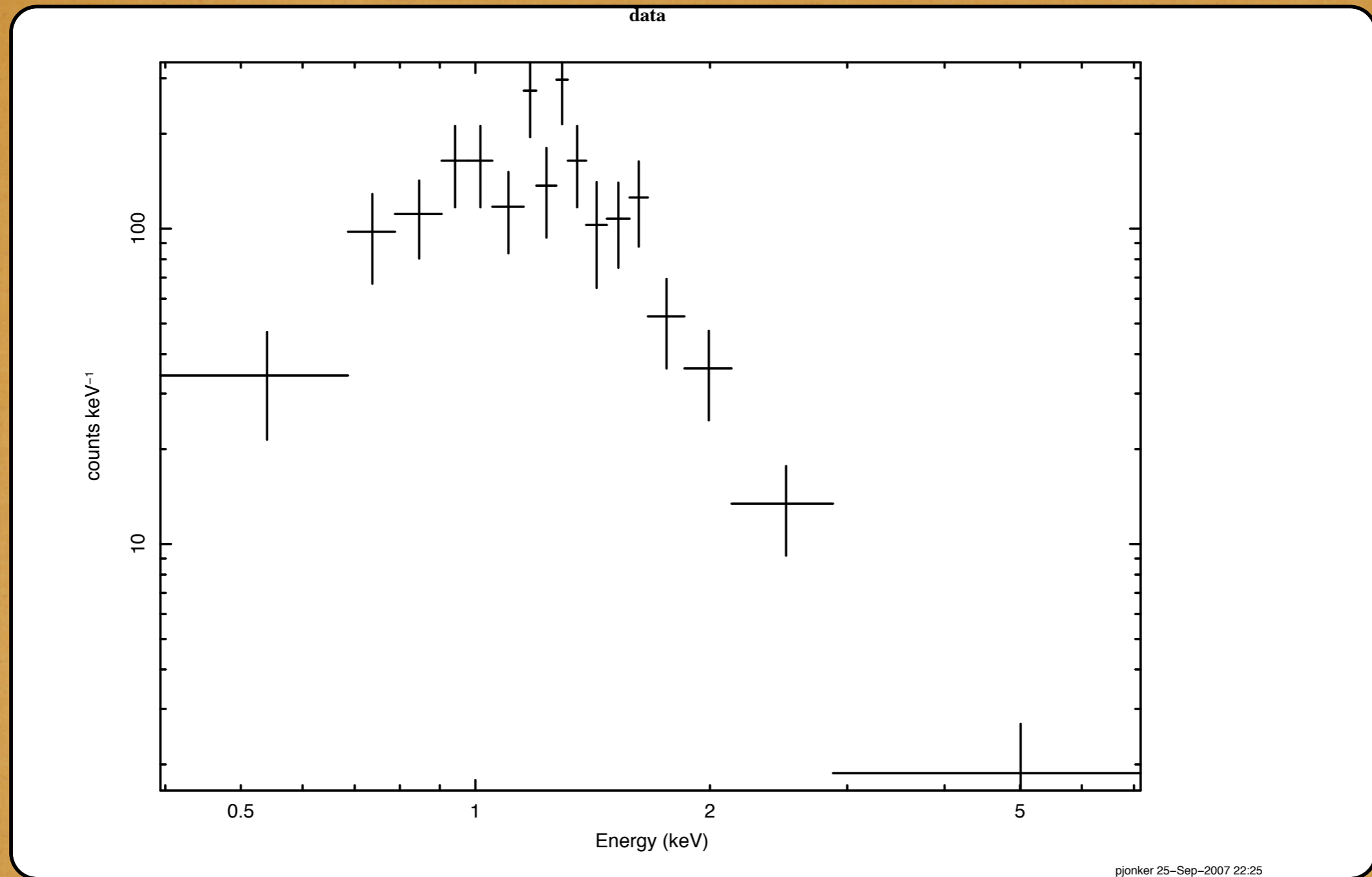


EXAMPLE FROM BEVINGTON & ROBERTSON 1992

CHANDRA CCD (ACIS) OBSERVATION OF AN X-RAY BINARY



SAME DATA AS BEFORE



GAUSSIAN APPROXIMATION FOR
ERRORS BUT AT LOW COUNTS GAUSS
AND POISSON ERRORS DIFFER

χ^2 FITTING PROVIDES:

BEST-FITTING PARAMETERS

AN ERROR ESTIMATE OF THE
UNCERTAINTY OF THE FITTED
PARAMETERS

A PROBABILITY THAT THE DATA IS
DRAWN FROM A PARENT POPULATION
DESCRIBED BY THE MODEL
PARAMETERS

FITTING A STRAIGHT LINE TO THE DATA

(SEE ALSO EXERCISE)

$$y_m(x_i, a, b) = a + bx_i$$

MINIMISE χ_i^2 TO FIND BEST-FITTING PARAMETERS

$$\frac{\partial \sum_{i=1}^N \left(\frac{y_i - a - bx_i}{\sigma_i} \right)^2}{\partial a} = 0 \quad \rightarrow$$

$$\sum \frac{y_i}{\sigma_i^2} - a \sum \frac{1}{\sigma_i^2} - b \sum \frac{x_i}{\sigma_i^2} = 0$$

&

$$\sum \frac{x_i y_i}{\sigma_i^2} - a \sum \frac{x_i}{\sigma_i^2} - b \sum \frac{x_i^2}{\sigma_i^2} = 0$$

$$a = \frac{\sum_i \frac{x_i^2}{\sigma_i^2} \sum_i \frac{y_i}{\sigma_i^2} - \sum_i \frac{x_i}{\sigma_i^2} \sum_i \frac{x_i y_i}{\sigma_i^2}}{\sum_i \frac{1}{\sigma_i^2} \sum_i \frac{x_i^2}{\sigma_i^2} - \left(\sum_i \frac{x_i}{\sigma_i^2} \right)^2}$$

DETERMINE ERRORS ON THE BEST-FITTING PARAMETERS

REMEMBER $\sigma_f^2 = \sigma_u^2 \left(\frac{\partial f}{\partial u} \right)^2 + \sigma_v^2 \left(\frac{\partial f}{\partial v} \right)^2 + \dots$

$$\sigma_a^2 = \sum_{i=1}^N \left[\sigma_i^2 \frac{\partial a}{\partial y_i} \right]^2$$

∂u & ∂v ETC ARE THE DIFFERENT MEASUREMENT VALUES y_i

$$\sigma_a^2 = \frac{\sum_i \frac{x_i^2}{\sigma_i^2}}{\sum_i \frac{1}{\sigma_i^2} \sum_i \frac{x_i^2}{\sigma_i^2} - \left(\sum_i \frac{x_i}{\sigma_i^2} \right)^2}$$

SIMILARLY

$$\sigma_b^2 = \frac{\sum_i \frac{1}{\sigma_i^2}}{\sum_i \frac{1}{\sigma_i^2} \sum_i \frac{x_i^2}{\sigma_i^2} - \left(\sum_i \frac{x_i}{\sigma_i^2} \right)^2}$$

FINALLY CALCULATE THE PROBABILITY
OF OBTAINING THE χ^2
BY CHANCE

$$P(\chi_{obs}^2) = \text{gammq}\left(\frac{N - M}{2}, \frac{\chi_{obs}^2}{2}\right)$$

FOR THE STRAIGHT LINE FIT $M=2$

$\nu = N - M$ DEGREES OF FREEDOM

REDUCED χ_ν^2

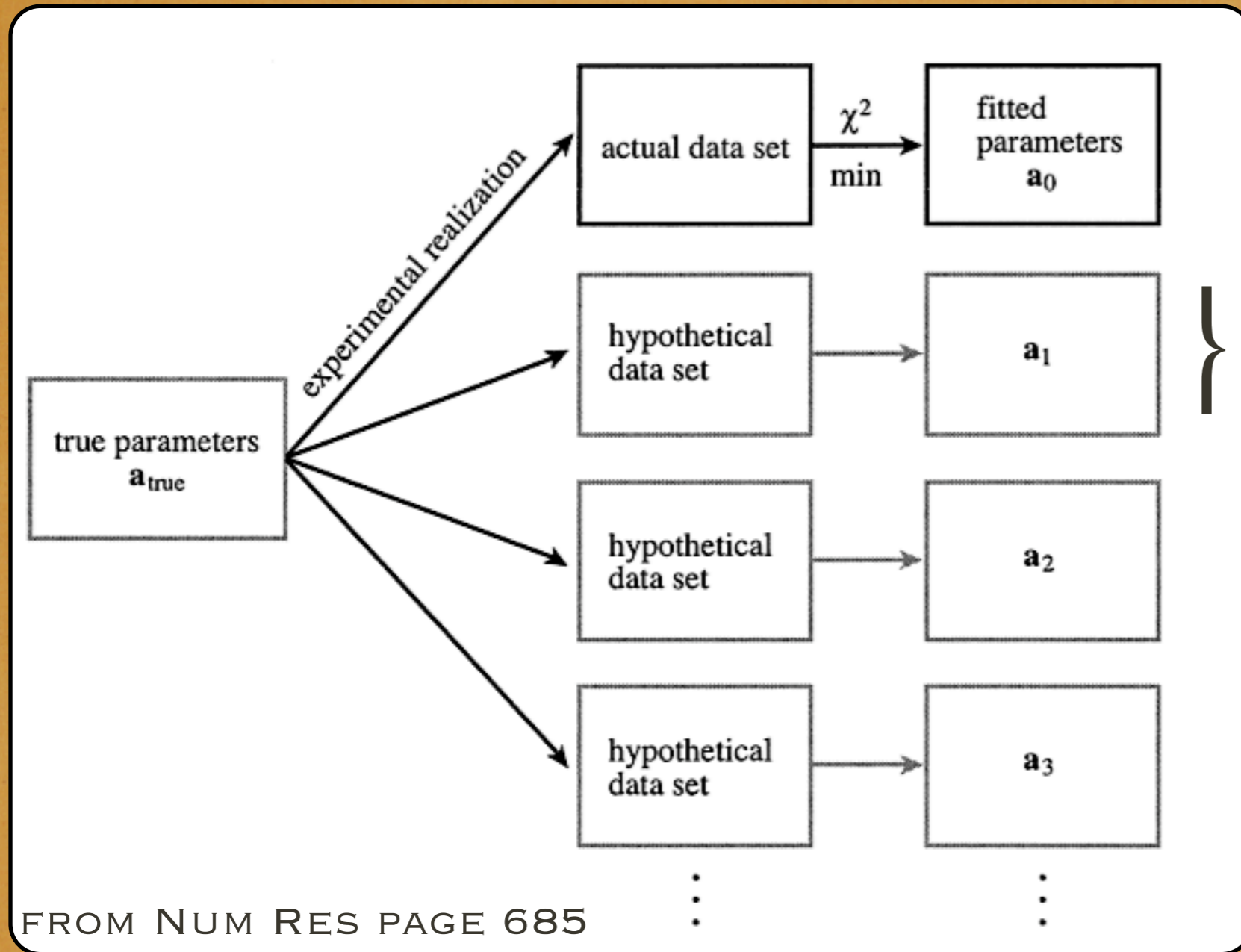
$$\chi_\nu^2 \equiv \frac{\chi^2}{\nu}$$

FOR DATA FITTING: $\chi_\nu^2 \sim 1$ $\chi^2 \approx \nu$

$$\sigma_{\chi^2} = \sqrt{2\nu}$$

ESTIMATING CONFIDENCE LIMITS

MONTE CARLO SIMULATIONS



\vec{a}_0

ONE DRAW FROM
THE DISTRIBUTION
OF A'S

FROM NUM RES PAGE 685

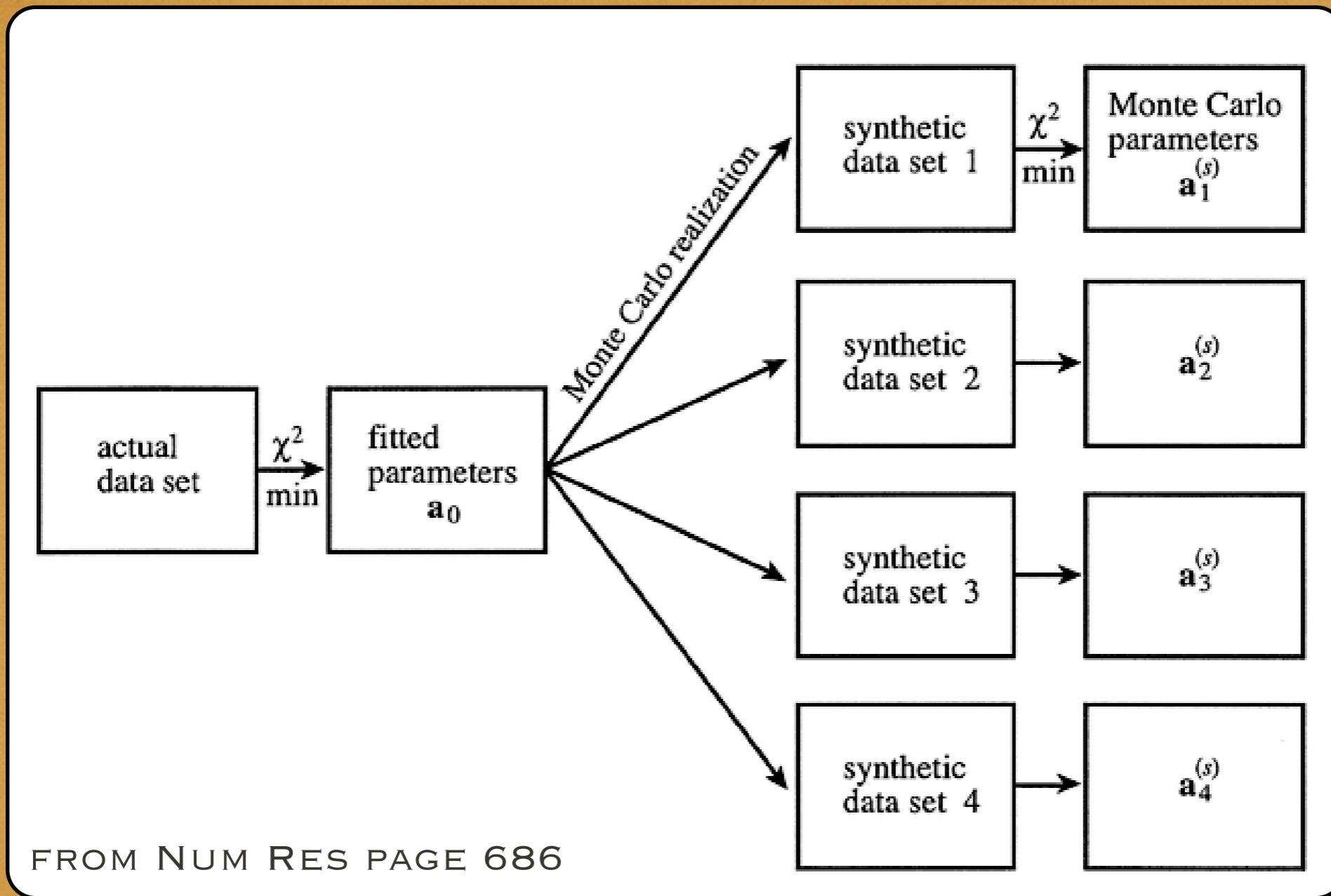
ASSUME THAT THE DISTRIBUTION OF

$$a_i - a_0$$

IS CLOSE TO THE PROBABILITY DISTRIBUTION

$$a_i - a_{true}$$

MONTE CARLO SIMULATIONS



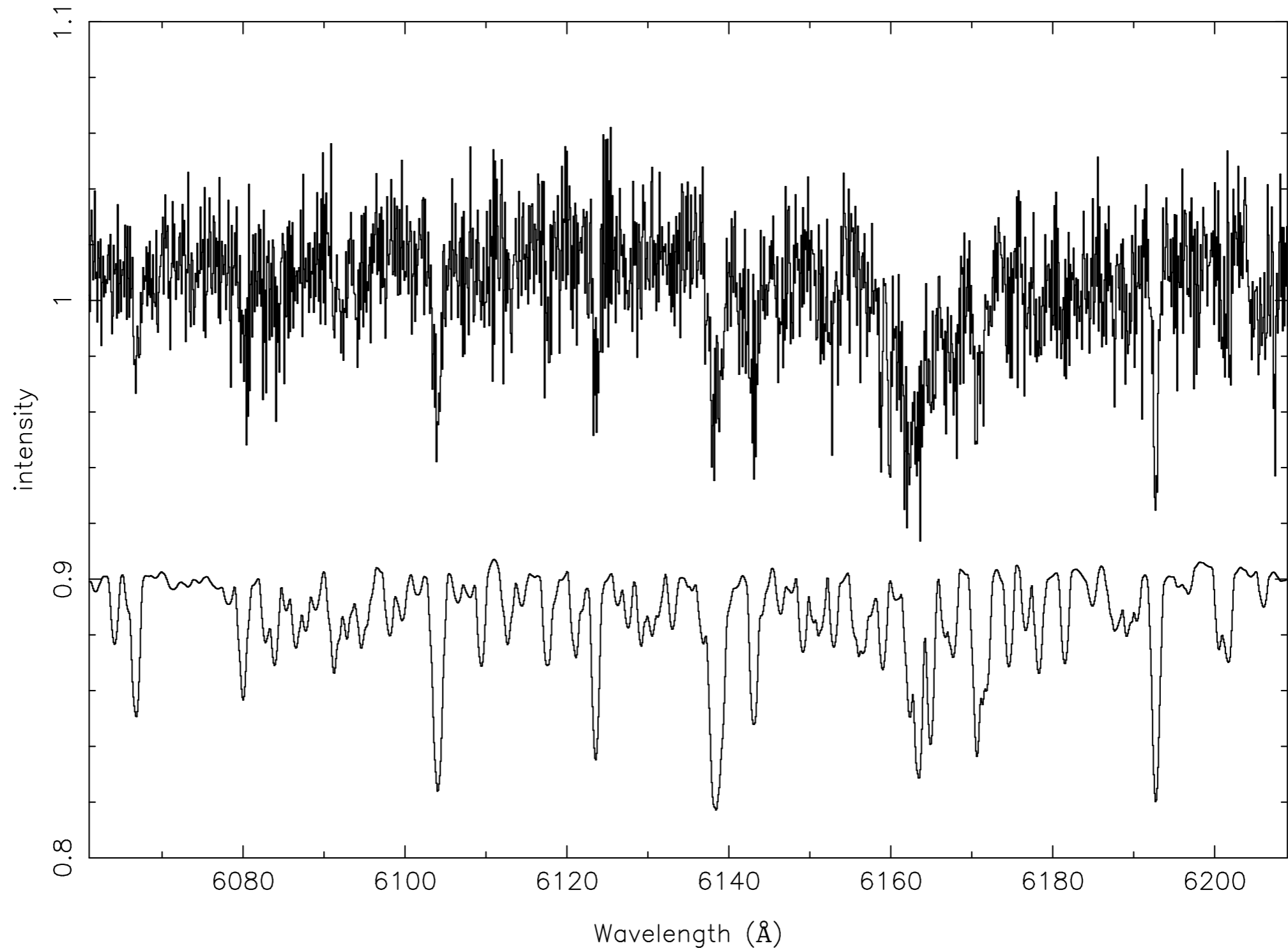
CALCULATE DISTRIBUTION OF $a_i - a_0$

BY SIMULATING MANY SETS OF DATA AND
USING χ^2 FITTING TO DETERMINE a_i

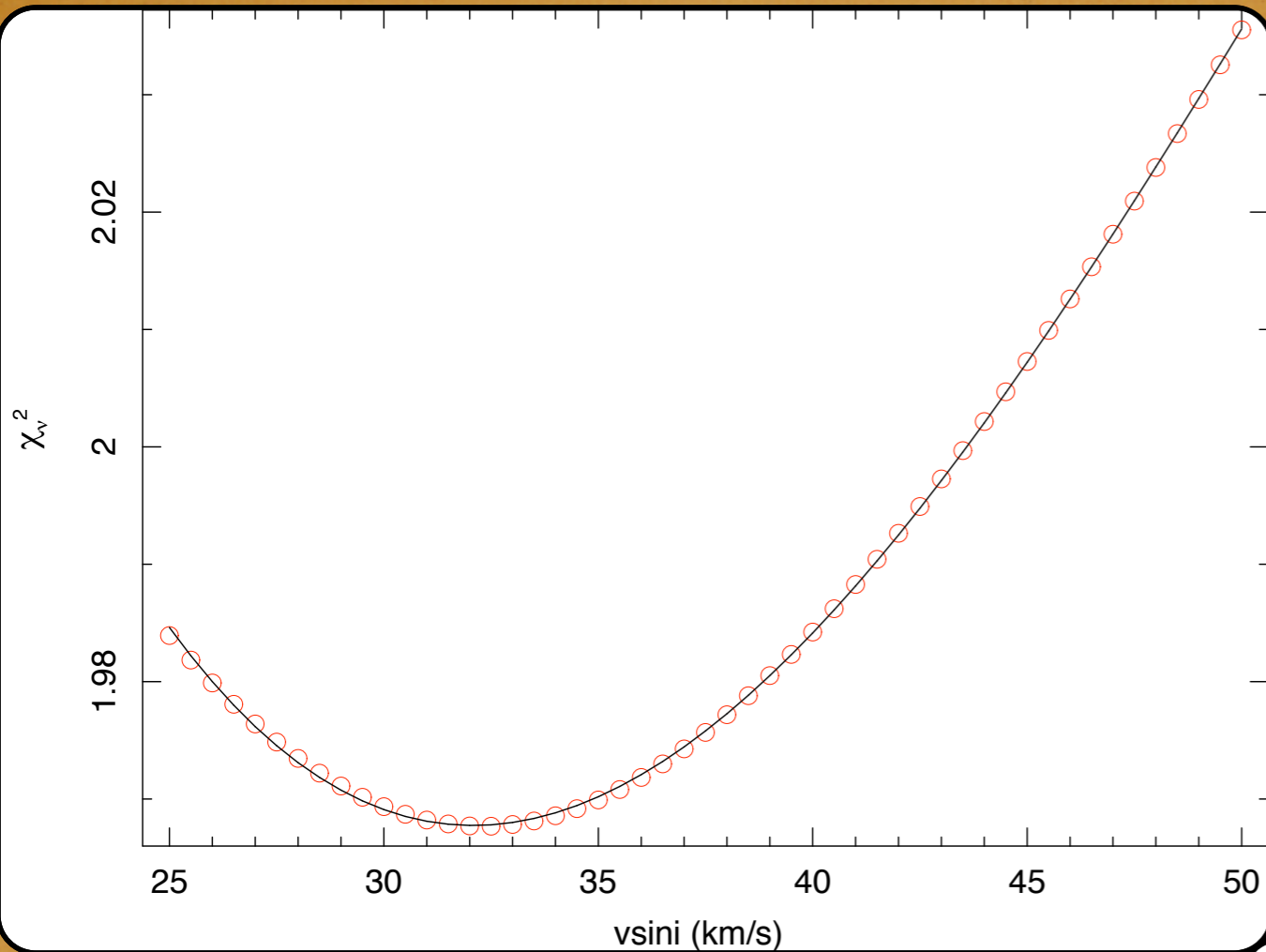
BOOTSTRAP METHOD AND APPLICATION

X-RAY BINARY V395 CAR

V395 Car / HD99322

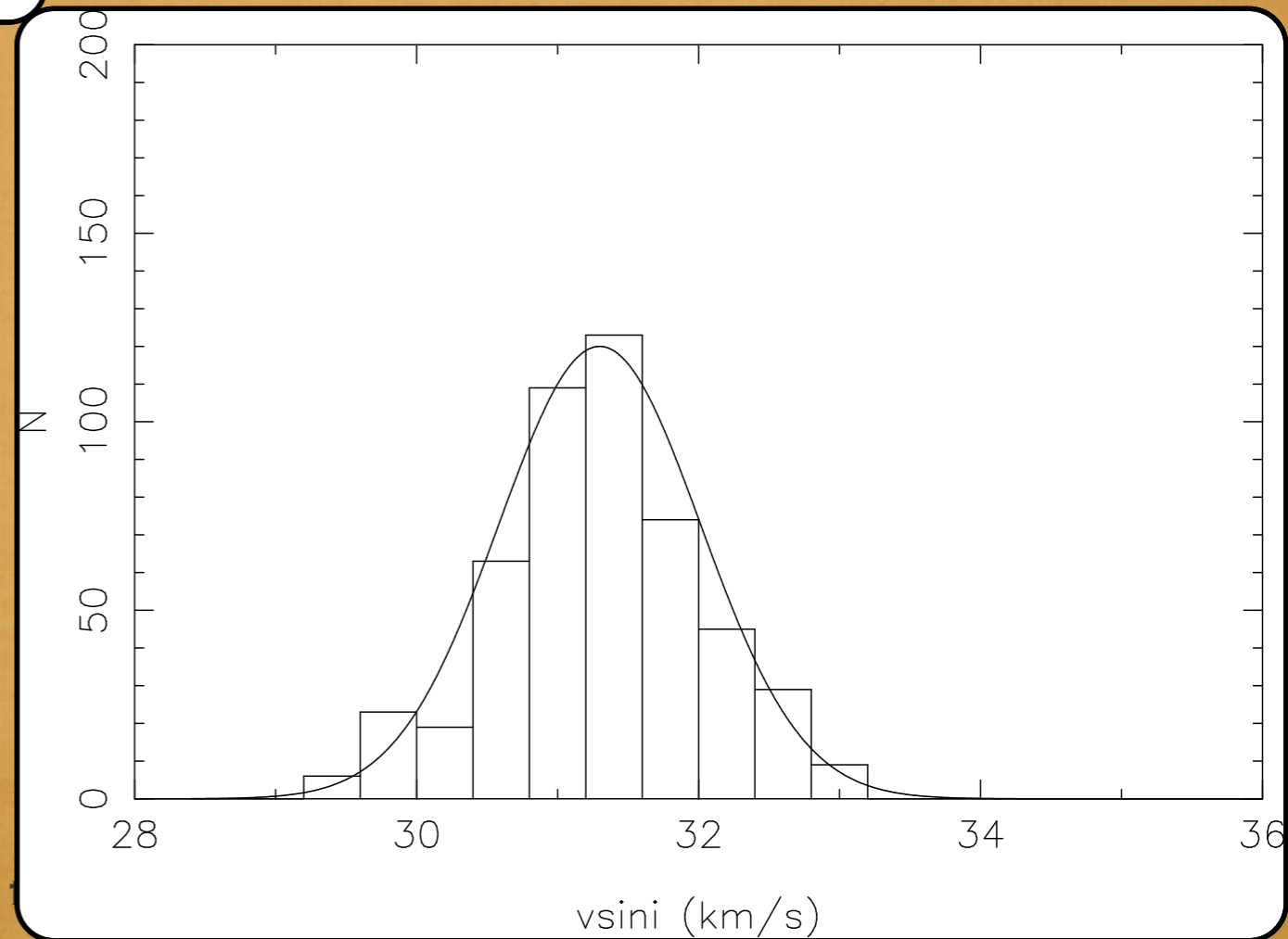


BROADENING AND OPTIMAL SUBTRACTION

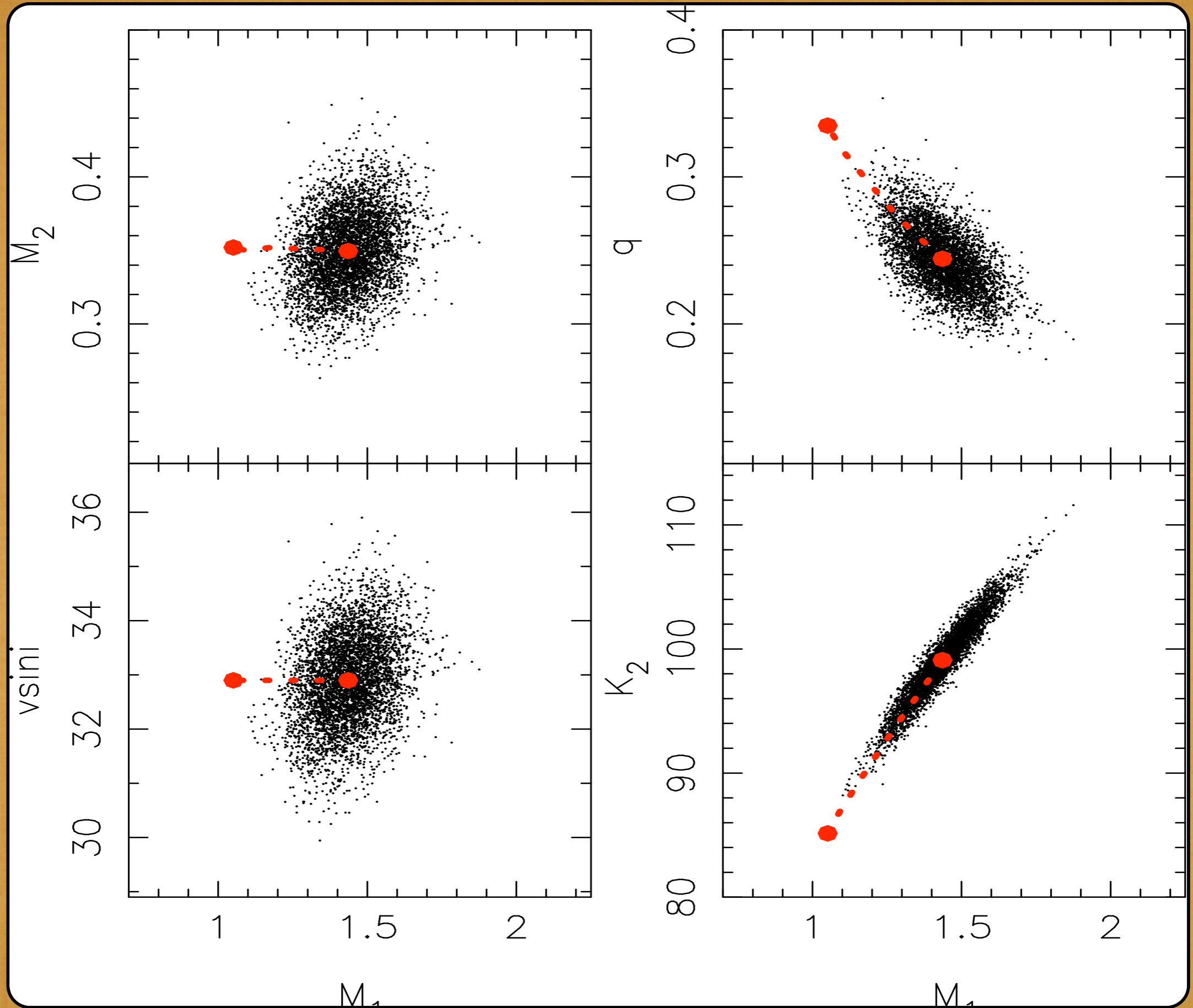


NOTE REDUCED χ_ν^2

BOOTSTRAP DETERMINED
ROTATIONAL VELOCITY

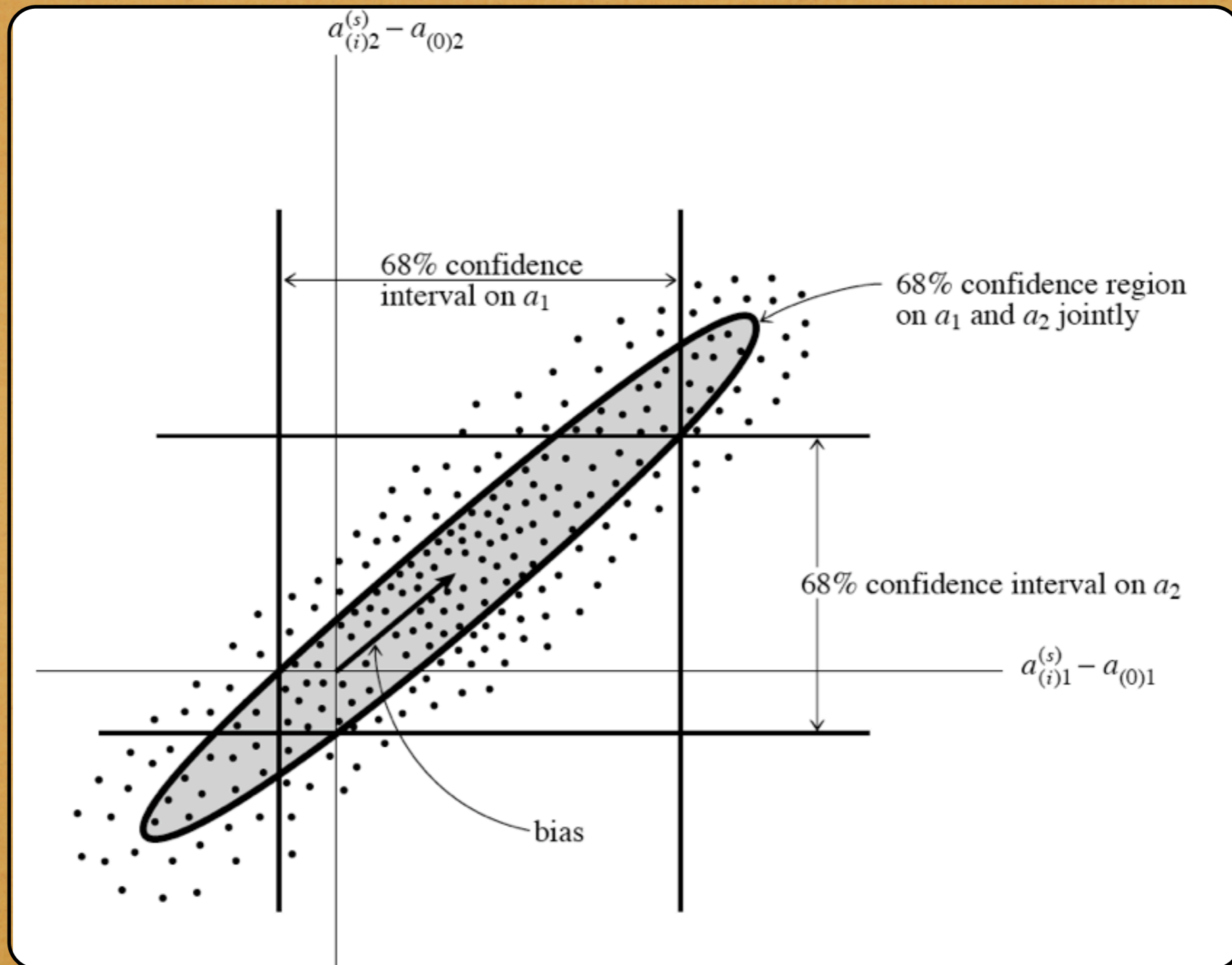


BOOTSTRAP METHOD: AN APPLICATION

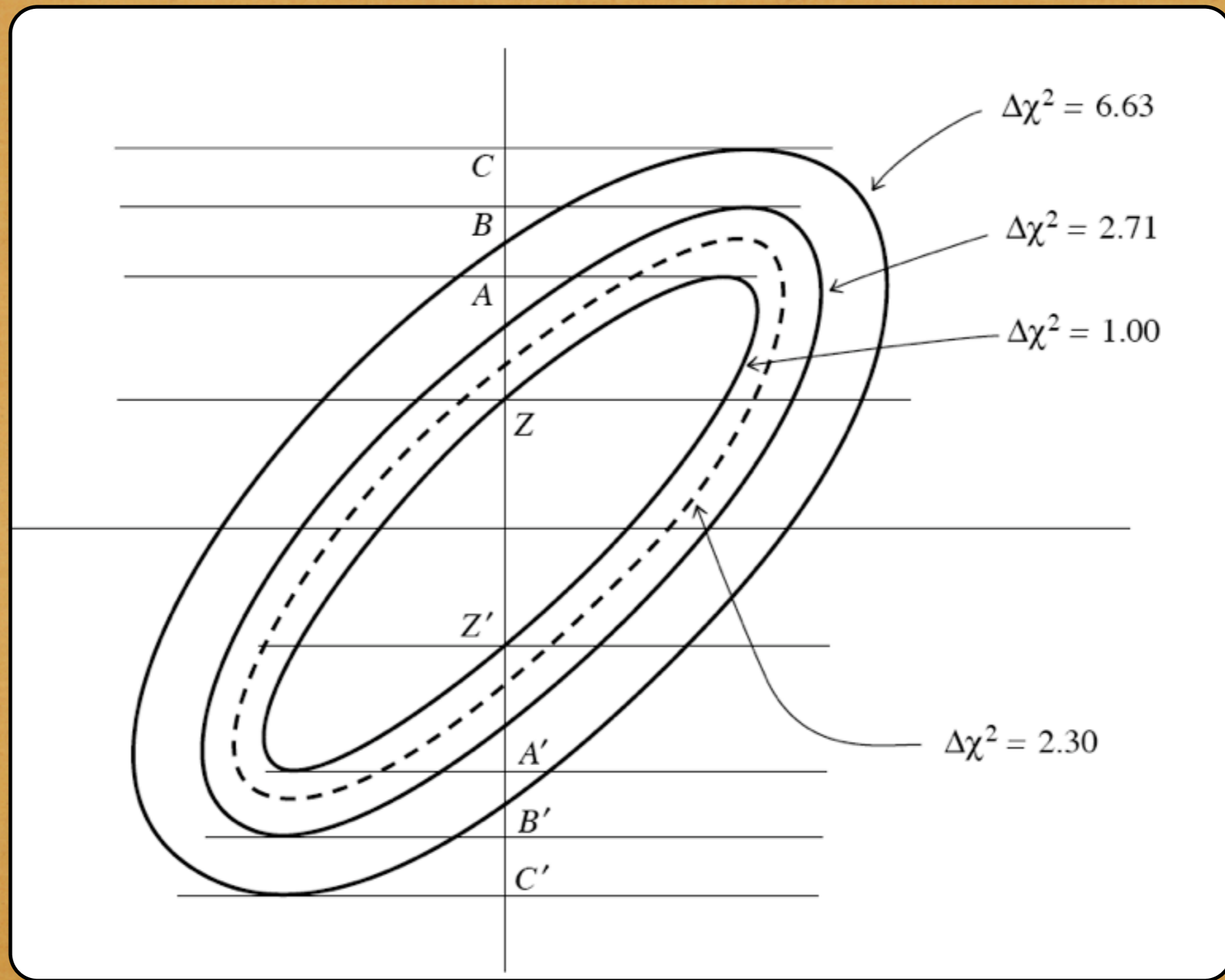


CONFIDENCE LIMITS

SINGLE VS. MULTIPLE PARAMETER CONFIDENCE REGION



PROJECTIONS



BE AWARE OF NON-GAUSSIAN DISTRIBUTIONS

